

Estimation of multivariate generalized gamma convolutions through Laguerre expansions

The beauty of multivariate Thorin classes

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Recall: Cumulant generating function: $K(\mathbf{t}) = \ln(\mathbb{E}(e^{\langle \mathbf{t}, \mathbf{X} \rangle}))$.

Definition (Multivariate Thorin Classes¹)

$$\mathbf{X} \sim \mathcal{G}_{d,n}(\boldsymbol{\alpha}, \mathbf{s}) \Leftrightarrow K(\mathbf{t}) = -\sum_{i=1}^n \alpha_i \ln(1 - \langle \mathbf{s}_i, \mathbf{t} \rangle)$$

$$\mathbf{X} \sim \mathcal{G}_d(\nu) \Leftrightarrow K(\mathbf{t}) = -\int \ln(1 - \langle \mathbf{s}, \mathbf{t} \rangle) \nu(d\mathbf{s}).$$

Property

\mathcal{G}_1 is closed w.r.t (independent) sums and products of random variables !

Example

All gammas, log-Normals, Paretos, α -stables, and sums and/or products of these are in \mathcal{G}_1 .

No estimation procedures currently available.

¹Lennart Bondesson. "On Univariate and Bivariate Generalized Gamma Convolutions". en. In: *Journal of Statistical Planning and Inference* 139.11 (Nov. 2009), pp. 3759–3765. ISSN: 03783758.

Motivation: interpretability of $\mathcal{G}_{d,n}$ models

$\mathbf{X} \in \mathcal{G}_{d,n}$ follows an additive risk-factor structure: there exists gamma random variables G_1, \dots, G_n such that:

$$\begin{pmatrix} X_1 \\ \dots \\ X_d \end{pmatrix} = \begin{pmatrix} s_{1,1} & \dots & \dots & s_{1,n} \\ \dots & \dots & \dots & \dots \\ s_{d,1} & \dots & \dots & s_{d,n} \end{pmatrix} \cdot \begin{pmatrix} G_1 \\ \dots \\ G_n \end{pmatrix}$$

(i) For all $i \in 1, \dots, n$, $G_i \sim \mathcal{G}_{1,1}(\alpha_i, 1)$

(ii) G_1, \dots, G_n are independent.

Since $s_{i,j}$ can always be zero, by increasing n (typically $n \gg d$) and using the infinite divisibility, we can approach any marginal in \mathcal{G}_1 , and we have a wide variety of dependence structures (asymmetry, tail dependency, etc.).

A generalized moment problem : Miles, Furman and Kuznetsov

Miles, Furman and Kuznetsov² exhibit in the univariate case a moment problem for the Thorin measure. A multivariate analogue of their train of thoughts might be the following:

For $\mathbf{X} \sim \mathcal{G}_d(\nu)$, $K(\mathbf{t}) = -\int \ln(1 - \langle \mathbf{s}, \mathbf{t} \rangle) \nu(\partial \mathbf{s})$. Derive \mathbf{i} times on \mathbf{t} to obtain:

$$\kappa_{\mathbf{i}, \mathbf{t}} = (|\mathbf{i}| - 1)! \int \frac{\mathbf{s}^{\mathbf{i}}}{(1 - \langle \mathbf{s}, \mathbf{t} \rangle)^{|\mathbf{i}|}} \nu(\partial \mathbf{s})$$

The substitution $\mathbf{x} = \frac{\mathbf{s}}{1 - \langle \mathbf{s}, \mathbf{t} \rangle} \iff \mathbf{s} = \frac{\mathbf{x}}{1 + \langle \mathbf{x}, \mathbf{t} \rangle}$ gives: $\xi_{\mathbf{i}} = \int_{\Delta_d(-\mathbf{t})} \mathbf{x}^{\mathbf{i}} \xi(\partial \mathbf{x})$, which can be seen as a d -variate generalized moment problem on the $(-\mathbf{t})$ -simplex.

Solving this moment problem for any \mathbf{t} (e.g. $-\mathbf{1}$) provides a \mathcal{G}_d parametrization corresponding to the shifted cumulants $\kappa_{\mathbf{i}, \mathbf{t}}$.

²Justin Miles, Edward Furman, and Alexey Kuznetsov. "Risk Aggregation: A General Approach via the Class of Generalized Gamma Convolutions". In: *Variance* (2019).

Unfortunately, no solution

Generalized moment problems have solvers through Lasserre's hierarchy of SDPs, see Helton & Nie³. **To even obtain a solution, the input data needs to be inside the cone of moments.**

- (i) The cumulants $\kappa_{i,t}$ needs to come from a true \mathcal{G}_d distribution so that ξ is in the cone of moments (300 digits of tanh-sinh integration per moment if $d = 1$ according to MFK, probably more as d increase)
- (ii) **For empirical cumulants, no way to define a coherent loss:** Should we take a Pareto front ? Give more weight to smaller moments ?

A least square approach could work for inexact cumulants, but we do not have a way to weight the objectives. **Also this breaks Lasserre's hierarchy simplicity..**

⇒ We need another approach.

³J. William Helton and Jiawang Nie. "A Semidefinite Approach for Truncated K-Moment Problems". en. In: *arXiv:1105.0410 [math]* (Sept. 2012), Jiawang Nie. "The A-Truncated K-Moment Problem". en. In: *arXiv:1210.6930 [math]* (Oct. 2012), Didier Henrion and Jérôme Malick. "Projection Methods in Conic Optimization". en. In: *Handbook on Semidefinite, Conic and Polynomial Optimization*. Ed. by Miguel F. Anjos and Jean B. Lasserre. Vol. 166. Boston, MA: Springer US, 2012, pp. 565–600. ISBN: 978-1-4614-0768-3 978-1-4614-0769-0.

The orthonormal Laguerre basis of $L_2(\mathbb{R}_+^d)$

Definition (Laguerre basis, see Comte⁴, Mabon⁵ and Dussap⁶)

For all $\mathbf{p} \in \mathbb{N}^d$, $\varphi_{\mathbf{p}}(\mathbf{x}) = \prod_{i=1}^d \varphi_{p_i}(x_i)$ where $\varphi_p(x) = \sqrt{2} \sum_{k=0}^p \binom{p}{k} \frac{(-2x)^k}{k!} e^{-x}$.

These functions form an orthonormal basis of $L_2(\mathbb{R}_+^d)$.

Therefore, every density f that is square-integrable can be expanded as :

$$f(\mathbf{x}) = \sum_{\mathbf{p} \in \mathbb{N}^d} a_{\mathbf{p}} \varphi_{\mathbf{p}}(\mathbf{x}) \text{ where } a_{\mathbf{p}} = \int \varphi_{\mathbf{p}}(\mathbf{x}) f(\mathbf{x}) d\mathbf{x}$$

Integrated square error loss: $L(\alpha, \mathbf{s}) = \sum_{k \leq m} (\hat{a}_k - a_k(\alpha, \mathbf{s}))^2$

⁴Fabienne Comte and Valentine Genon-Catalot. "Adaptive Laguerre Density Estimation for Mixed Poisson Models". en. In: *Electronic Journal of Statistics* 9.1 (2015), pp. 1113–1149. ISSN: 1935-7524.

⁵Gwenn lle Mabon. "Adaptive Deconvolution on the Non-Negative Real Line: Adaptive Deconvolution on \mathbb{R}_+ ". en. In: *Scandinavian Journal of Statistics* 44.3 (Sept. 2017), pp. 707–740. ISSN: 03036898.

⁶Florian Dussap. "Anisotropic Multivariate Deconvolution Using Projection on the Laguerre Basis". In: (2020).

Computing densities of $\mathcal{G}_{d,n}$ models: $(\alpha, s) \mapsto \kappa \longleftrightarrow \mu \longleftrightarrow a$

Algorithm 1: Laguerre coefficients of $\mathcal{G}_{d,n}(\alpha, s)$ random vectors

Input: Shapes $\alpha \in \mathbb{R}_+^d$, scales $s \in \mathcal{M}_{n,d}(\mathbb{R}_+)$, and truncation threshold $m \in \mathbb{N}^d$

Result: Laguerre coefficients $(a_k)_{k \leq m}$ of the $\mathcal{G}_{d,n}(\alpha, s)$ density

Compute the simplex version of the scales $x_i = \frac{s_i}{1+|s_i|}$ for all $i \in 1, \dots, n$.

Let $\kappa_0 = -\sum_{i=1}^n \alpha_i \ln(1 - |x_i|)$ and $a_0 = \mu_0 = \exp(\kappa_0)$

foreach $0 \neq k \leq m$ **do**

 Let $a_k = \mu_k = 0$, d be the index of the first k_i that is non-zero, $p = k$ and set $p_d = p_d - 1$.

 Let $\kappa_k = (|k| - 1)! \sum_{i=1}^n \alpha_i x_i^{k_i}$

foreach $l \leq p$ **do**

 Set $\mu_k += (\mu_l) (\kappa_{k-l}) \binom{p}{l}$ according to efficient Faà di Bruno's algorithm from Miatto^a

 Set $a_k += \mu_l \binom{k}{l} \frac{(-2)^{|l|}}{l!}$

end

 Set $a_k += \mu_k \frac{(-2)^{|k|}}{k!}$

end

$a = \sqrt{2}^d a$

Return a

^aFilippo M. Miatto. "Recursive Multivariate Derivatives of $\mathbb{S}e^{\{f(X_1, \dots, X_n)\}}$ of Arbitrary Order". en. In: *arXiv:1911.11722 [cs, math]* (Nov. 2019).

Analysis of the loss: The ε -w.b. condition

Definition (ε -well-behaved $\mathcal{G}_{d,n}(\alpha, \mathbf{s})$)

$\mathcal{G}_{1,n}(\alpha, \mathbf{s})$ is ε -w.b. $\iff |\alpha| > 1$ and $\mathbf{s} \in]\frac{\varepsilon}{2+\varepsilon}, \frac{2+\varepsilon}{\varepsilon}[^n$.

$\mathcal{G}_{d,n}(\alpha, \mathbf{s})$ is ε -w.b. \iff [Technical, useful but boring definition]

Property (well-behaved $\mathcal{G}_{d,n}(\alpha, \mathbf{s})$)

$\mathcal{G}_{d,n}(\alpha, \mathbf{s})$ is w.b. $\iff |\alpha| > 1$ and $\forall I \subseteq \{1, \dots, n\}$ such that $\sum_{i \in I} \alpha_i > \sum_{i \notin I} \alpha_i$, $\text{Ker}(\mathbf{s}_I) = \{\mathbf{0}\}$.

Example (Simple w.b. examples)

As soon as $|\nu| = |\alpha| > 1$, the following are w.b. :

- (i) All univariate gamma convolution
- (ii) All $\mathcal{G}_{d,n}$ with independent marginals and all invertible linear transformation of them.
- (iii) All finite convolution of well-behaved gamma convolutions.

Property (Exponential decay of Laguerre coefficients)

For any $\varepsilon' > 0$ and any dimension d , $\exists B(d, \varepsilon') \in]0, +\infty[$, such that Laguerre coefficients $(a_{\mathbf{k}})_{\mathbf{k} \in \mathbb{N}^d}$ of any d -variate ε -well-behaved gamma convolution, $\varepsilon > \varepsilon'$, verify:

$$|a_{\mathbf{k}}| \leq B(d, \varepsilon')(1 + \varepsilon')^{-|\mathbf{k}|}.$$

Implication: The ISE loss $L(\alpha, \mathbf{s}) = \sum_{\mathbf{k} \leq m} (\hat{a}_{\mathbf{k}} - a_{\mathbf{k}}(\alpha, \mathbf{s}))^2$ produces consistent estimators of (α, \mathbf{s}) .

Proofs leverage analytics combinatorics in several variables on the function $R = \exp \circ K \circ \mathbf{h}$, which happens to be the generating function of Laguerre coefficients, where \mathbf{h} is a componentwise Möbius transform.

Remark on implementation and complexity

We need to minimize a quadratic loss on the coefficients computed through Algorithm 1.

Therefore, our loss is:

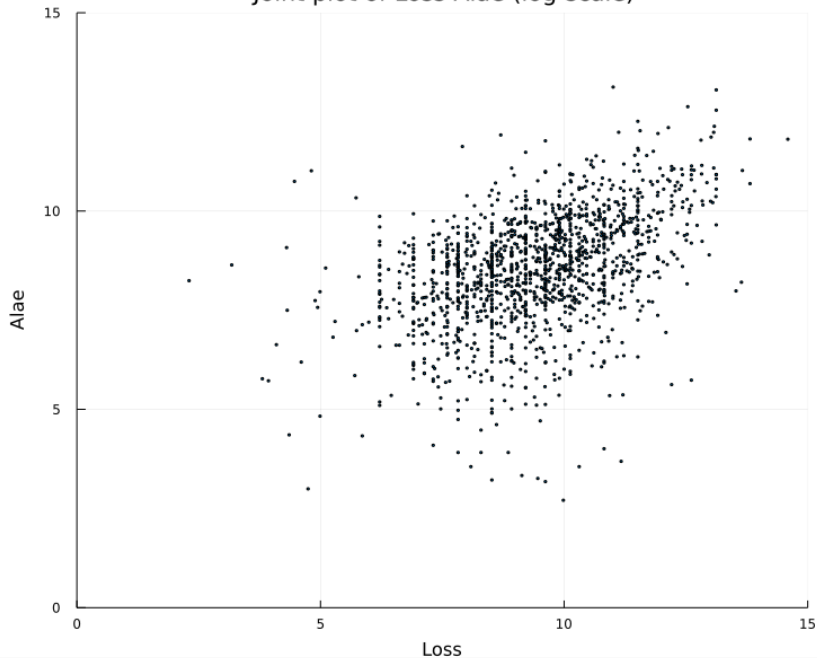
Combinatorial	⇒	Arbitrary precision	
Highly recursive	⇒	Compiled code	⇒ Julia
Highly non-convex	⇒	Global optimization	

Our package `ThorinDistributions.jl`⁷, available on `github`, implements this loss efficiently.

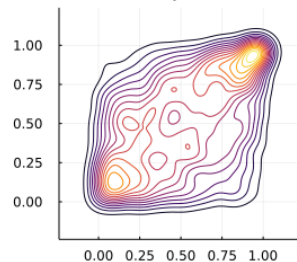
⁷Oskar Laverny. *lrv/ThorinDistributions.jl: ThorinDistributions.jl v0.1*. Version v0.1. Mar. 2021. DOI: 10.5281/zenodo.4644109. URL: <https://doi.org/10.5281/zenodo.4644109>.

Crunching numbers on Loss-Alae: the dataset

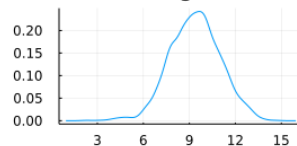
Joint plot of Loss-Alae (log-scale)



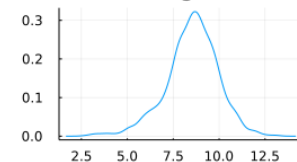
Copula



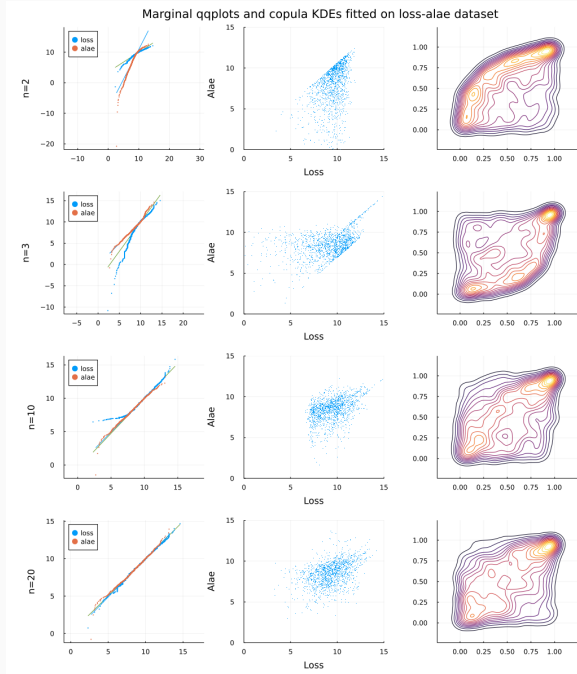
Loss (log-scale)



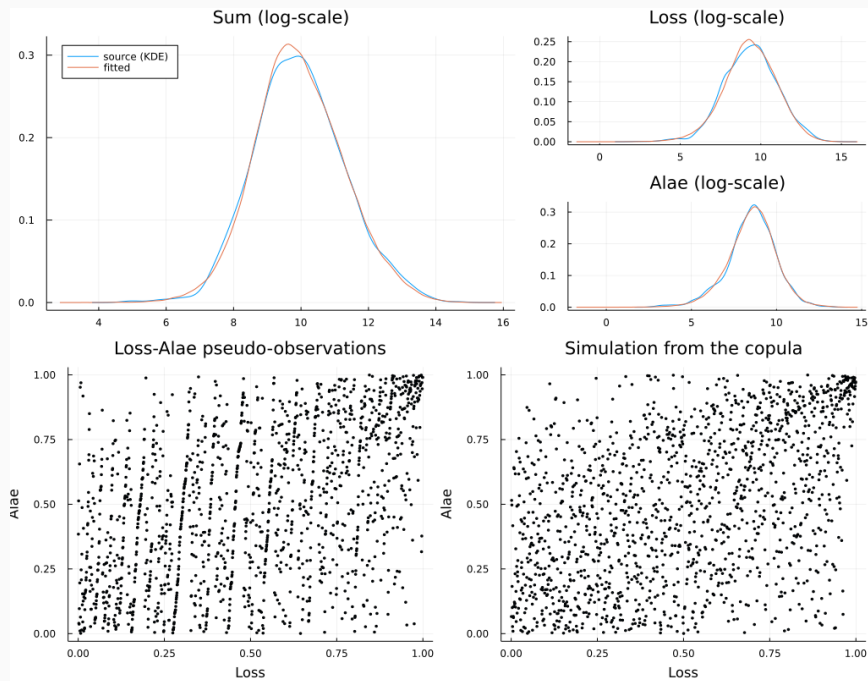
Alae (log-scale)



Crunching numbers on Loss-Alae: Results for several numbers of gammas



Crunching numbers on Loss-Alae: Results for $n = 20$



Drawbacks of the approach and potential solutions

The total number of coefficients to compute is $\prod_{i=1}^d (m_i + 1)$. This is quite unpractical when d gets large, and therefore restricts the use of the algorithm to $d < 5$ or 6 to have a performant procedure.

However, there is a way out of the curse: If $X_i \sim \mathcal{G}_{1, n_i}(\alpha_i, \mathbf{s}_i)$, denoting

$\mathbb{N}_n^d = \{\mathbf{p} \in \mathbb{N}^d : \forall i, 0 \leq p_i \leq n_i\}$ and $N = \prod_{i=1}^d (n_i + 1)$ its cardinal, any multivariate gamma convolutions with these marginals is a $\{\mathcal{G}_{d, N}(\mathbf{a}, \mathbf{S})\}$ such that:

- (i) $\mathbf{S} = \{(s_{1, i_1}, \dots, s_{d, i_d}) : \forall \mathbf{i} \in \mathbb{N}_n^d\}$, with $s_{i, 0} = 0$ for all i .
- (ii) $\mathbf{a} \in \mathbb{R}_+^{\mathbb{N}_n^d}$ and for all i, j , $\sum_{\mathbf{p} \in \mathbb{N}_n^d} a_{\mathbf{p}} \mathbf{1}_{p_i=j} = \alpha_{i, j}$. (this constraint is linear).

Last, for a given constant \mathbf{c} to be chosen, $\langle \mathbf{c}, \mathbf{X} \rangle \sim \mathcal{G}_{1, N}(\mathbf{a}, \{\langle \mathbf{c}, \mathbf{s} \rangle : \mathbf{s} \in \mathbf{S}\})$.

Estimating $\langle \mathbf{c}, \mathbf{X} \rangle$ conditionally on the marginals is therefore enough.

- (i) The univariate Thorin class is wide: LN, Pareto, α -stable, (some) Weibull, ...
- (ii) The Multivariate analogue provides an asymmetrical dependence structure which include tail dependency, and can take a lot of different shapes.
- (iii) Deconvolution is a hard inverse problem, and estimation of these distributions is complicated.
- (iv) The final additive risk-factor model gives easy interpretation of parameters and easy aggregation schemes.
- (v) A better approach for high dimensional cases might be possible (working on it).

Details and other illustrated applications are available in our paper⁸ and Julia package repo⁹.

Thanks !

⁸Oskar Laverny, Esterina Masiello, Véronique Maume-Deschamps, and Didier Rullière. *Estimation of multivariate generalized gamma convolutions through Laguerre expansions*. 2021. arXiv: 2103.03200 [math.ST].

⁹Oskar Laverny. *Irvv/ThorinDistributions.jl: ThorinDistributions.jl v0.1*. Version v0.1. Mar. 2021. DOI: 10.5281/zenodo.4644109. URL: <https://doi.org/10.5281/zenodo.4644109>.